Television viewing and risk of mortality: exploring the biological plausibility

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**Abstract**

**Background and aims:** Television (TV) viewing is a major component of leisure sedentary time, and has been consistently associated with cardiovascular disease. We examined the extent to which metabolic biomarkers explain the association between TV viewing and mortality.

**Methods:** Participants (N=8,451, aged 64.8±9.9 yrs) were drawn from The English Longitudinal Study of Ageing (ELSA), a national prospective cohort study of community-dwelling men and women living in England. The individual participant data were linked with death records from the National Health Service registries from 2008 to 2012. Cox proportional hazards models were used to estimate the risk of death according to time spent watching TV, with biomarkers added in a step wise fashion to estimate potential mediation.

**Results:** Over an average follow up of 4 years (33,832 person years) there were 370 deaths. In models adjusted for comorbidities, psychosocial factors, and health behaviours including physical activity, there was an association between TV viewing and mortality (≥6 hrs per day vs. < 2 hrs per day [Ref]; Hazard ratio = 1.98, 95% CI, 1.25, 3.15). Adjustment for inflammatory markers (C-reactive protein and fibrinogen) accounted for ~15.7% of the association between TV viewing and mortality, but metabolic risk factors (HDL-cholesterol, triglycerides, glycated haemoglobin) did not contribute.

**Conclusion:** The association between TV viewing and mortality was partly mediated by inflammatory markers, although the relationship remains largely unexplained.

**Key words:** Sedentary; Ageing; Mortality; Inflammation; Biomarkers
Introduction

Television (TV) viewing is a major component of leisure sedentary time, and has been consistently associated with cardiovascular disease (CVD) and mortality in population cohort studies.\textsuperscript{1-5} Sedentary behaviour has been conceptualised as a distinct domain of behaviour, which may pose a risk to health independently of physical activity. High levels of physical activity appear to partly attenuate associations between TV time and mortality,\textsuperscript{1} but it is unclear whether the underlying mechanisms of this association are different than those described in the physical activity literature. Inflammatory and haemostatic risk factors, and blood pressure appear to explain the greatest proportion of the inverse association between physical activity and CVD events,\textsuperscript{6,7} but the mechanisms explaining the association between TV viewing and mortality remain poorly described.

Previous studies have explored associations of TV viewing with changes in inflammatory and metabolic risk factors although results have often been inconsistent,\textsuperscript{8-13} and these studies did not include follow-up of mortality to examine mediation effects. The main aim of this study was to examine the extent to which inflammatory and metabolic biomarkers explain the association between TV viewing and mortality in a sample of community dwelling older men and women.
Materials and Methods

Study sample and procedures

The English Longitudinal Study of Ageing (ELSA) is a cohort study of older adults previously described.\textsuperscript{14} Data on TV viewing was first collected at wave 4 (2008/9), and was thus used as the baseline for the present analyses. The individual participant data were linked with death records from National Health Service registries for all consenting respondents (96.5% of the sample) up to February 2012, and primary cause of death were based on the International Classification of Diseases, Ninth (ICD-9) and Tenth (ICD-10) Revisions. Codes corresponding to CVD mortality were 390-459 for ICD-9 and I01-I99 for ICD-10 and cancer mortality were 140-239 for ICD-9 and C00-D48 for ICD-10. Participants with clotting and bleeding disorders, or taking anti-coagulant medication were excluded from providing blood. Participants gave full informed consent to participate in the study. Ethical approval was obtained from the London Multi-center Research Ethics Committee, compliant with the Declaration of Helsinki.

Baseline TV viewing and physical activity

Television viewing was assessed from self report as previously described.\textsuperscript{11, 15} The measure of TV viewing time was validated in 105 (55 men, 50 women) ELSA participants (aged 66.3±6.5 yrs) using objective accelerometry devices (unpublished data). Participants wore a GeneActiv device on their wrist for seven consecutive days. The categorical TV exposure variable described above was weakly correlated with objectively assessed hours per day of sedentary time ($r = 0.24$, $p=0.014$), which might reflect the fact TV only represents a proportion of overall sedentary time. The questions on TV viewing were administered two years later in the whole cohort and moderate reliability was observed (Pearson’s $r = 0.55$, $p<0.001$).
Participants were asked how often (> once a week, once a week, 1 – 3 times a month, and hardly ever/never) they took part in vigorous, moderate- and low-intensity physical activity. Prompt cards were used to assist participants in the interpretation of physical activity intensities. The physical activity and TV viewing measures used in ELSA have demonstrated convergent validity in relation to psychosocial, physical and biochemical risk factors in previous work.  

Clinical assessments
Nurses collected anthropometric data (weight, height), and blood samples. Body weight was measured without shoes and in light clothing using Tanita electronic scales, and height was measured using a Stadiometer with the Frankfort plane in the horizontal position. Body mass index (BMI) was calculated using the standard formula [weight (kilograms)/height (meters) squared]. Systolic and diastolic blood pressure was measured with an Omron HEM-907 monitor and an average of the second and third recordings was used for the present analyses. Blood samples were analyzed for several biomarkers, including high sensitivity C-reactive protein [CRP], fibrinogen, HDL-cholesterol, triglycerides, glycated haemoglobin. The samples were analysed at the Royal Victoria Infirmary (Newcastle-upon-Tyne, UK), described elsewhere.

Covariates
Trained interviewers asked questions on cigarette smoking (current, ex-smoker or non-smoker), alcohol intake (categorised as; at least five times a week, at least once a week, monthly, rarely/never), chronic illnesses (at least one vs. none), and depressive symptoms (using the 8-item Centre of Epidemiological Studies Depression scale). Disability was
denoted as reporting difficulties in one or more activities from questions on perceived difficulties in six basic activities (e.g. difficulty dressing) and seven instrumental activities of daily living (e.g. difficulty preparing a hot meal).21

Statistical analysis

We used Cox proportional hazards regression models to examine associations between TV viewing and death. Years were the time scale for the follow-up calculated from age at death, and for participants with no record of an event, the data were censored at February 2012. Plots of the Nelson-Aalen cumulative hazard estimates were examined to assess the proportional hazards assumption. The exposure variable was treated as a categorical variable in the main analyses but models were also run with TV viewing as a continuous variable using the standardised score (1 SD=4.2 hrs/d TV viewing). We estimated models that initially contained the covariates age and sex (model 1), with further adjustment for physical activity, smoking, alcohol intake, depressive symptoms (CES-D>3), long standing illness, disability (model 2). These covariates were considered as confounders and selected a priori based on previous literature.1,11,17 Sensitivity analyses were conducted removing deaths in the first year of follow up. In order to test for biological mediation we separately added inflammatory markers (log CRP, fibrinogen) and metabolic risk factors (HDL-cholesterol, triglycerides, glycated haemoglobin) into the models, all as continuous variables. BMI (as a categorical variable; <25; 25 – 29.99; ≥30 kg/m²) was added as a covariate in these models as it remains unclear if adiposity lies on the causal pathway from TV viewing to mortality or acts as a predisposing factor to sedentary behaviour.22 Mediation
was calculated from percentage changes in coefficients (In HR). All analyses were conducted using SPSS version 22 (SPSS, Chicago, IL).

**Results**

A sample size of 9,126 provided complete data on time spent TV viewing, although 357 participants did not provide consent to mortality linkage and a further 318 had missing data on other covariates leaving a final analytic sample of 8,451. There were no differences between participants excluded and the analytic sample. Biomarker data were available in a sub-sample of participants that completed the health examination (n=5,033).

The baseline characteristics of the sample are shown in Table 1. High levels of TV viewing (>6 hr/d) were reported in 29.1% of participants. Participants viewing high levels of TV were more likely to be female, smokers, undertake little vigorous exercise, report chronic illness, depression, and disability. Participants in the highest TV viewing category were also more likely to be obese and display a less favourable biomarker profile.

Over an average follow up of 4 years there were 370 deaths (78 CVD; 91 cancer). In models adjusted for comorbidities, psychosocial factors, and health behaviours including physical activity, there was an association between TV viewing and mortality (see Table 2). When the exposure data were treated continuously there was an association between TV (per SD unit increase) and mortality (fully adjusted HR= 1.17, 95% CI, 1.06, 1.28). In sensitivity analyses the results remained unchanged after removal of 72 deaths in first year of follow up (Table 2; Model 3). In analyses examining cause specific deaths the point estimates for CVD (fully adjusted HR per SD unit = 1.22, 95% CI, 1.00, 1.49) and cancer (fully adjusted HR per unit =
1.16, 95% CI, 0.96, 1.39) were similar to those of all-cause mortality albeit with wider confidence intervals owing to lower numbers of deaths.

We examined associations between biomarkers and mortality (Table 3). In models mutually adjusted for all biomarkers and BMI, fibrinogen (p<0.001) was most strongly predictive of mortality, whilst HDL-cholesterol (p=0.087) and CRP (p=0.061) reached marginal significance. Biological mediation was examined by separately adding inflammatory markers (CRP, fibrinogen) and metabolic risk factors (HDL-cholesterol, triglycerides, glycated haemoglobin) into the models of TV viewing and mortality (Table 4). The addition of inflammatory markers (C-reactive protein and fibrinogen) into the models accounted for ~15.7% of the association between TV viewing and mortality, although no attenuation in effect estimates were observed when metabolic risk factors (HDL-cholesterol, triglycerides, glycated haemoglobin) were added. We repeated these analyses using CVD death as the outcome and results were similar (Table 4).

Discussion

The main aim was to examine the extent to which biological risk factors mediate the association between TV viewing and mortality. We replicated previous studies by demonstrating an association between TV viewing and mortality. We further demonstrated the novel finding that this association was partly mediated by inflammatory markers. Demonstrating biological plausibility is important to understand the causal nature of an association. However, the association between TV viewing and mortality remained largely unexplained using the biomarkers available in this study.
Sedentary behaviours such as TV viewing are characterised by a lack of muscle contraction. Muscle tissue expresses various transcriptional co-activators such as PGC1α, which lead to a transient release of myokines (during exercise) that may promote systemic anti-inflammatory effects. Research has demonstrated that prolonged sitting is associated with the expression of various genes linked to inflammatory responses. Fibrinogen, the strongest predictor of mortality in the present study, may also have relevance in terms of elevated risk of vascular conditions, particularly venous thrombosis. Recent experimental data demonstrated increases in plasma fibrinogen with prolonged uninterrupted sitting that was attenuated with active breaks. Excess TV viewing may reflect chronic energy imbalance indicated by excess adiposity and metabolic disorders, although in the present study classic metabolic risk factors did not explain the link between TV viewing and mortality. Recent evidence has suggested that physical inactivity may be a consequence of obesity, which may explain why the associated metabolic disorders are not on the causal pathways linking sedentary behaviours with disease outcomes. Since conventional risk factors appear to loose predictive power in older populations, novel biomarkers may have been more relevant in the present cohort and explained a greater proportion of association between TV and mortality.

TV viewing was measured at the same time point as the biomarkers, thus we cannot establish the precise nature of the temporal relationship between these factors. Nevertheless, previously published data from ELSA has demonstrated a prospective association between TV viewing and change in inflammatory markers. Our analysis
supports the main criteria for mediation, that is, the key exposure (TV viewing) was associated with biomarkers (potential mediators), the potential mediators predicted outcome (mortality), and exposure predicted outcome; this association was attenuated after adjustment for the mediators (inflammatory markers).

Our study has some notable strengths, including the use of a large national sample of community-dwelling men and women; the inclusion of a wide range of potential biological mediators; and adjustment for important confounding factors. There are also some limitations. Our measure of sedentary behavior was limited to leisure TV viewing thus our results cannot be generalized to total sedentary time or sitting. Albeit in this older sample of adults it is likely that TV viewing is still by far the most predominant screen based activity. Our measure of sedentary behavior was limited to leisure TV viewing thus our results cannot be generalized to total sedentary time or sitting. Albeit in this older sample of adults it is likely that TV viewing is still by far the most predominant screen based activity. Nevertheless, our TV exposure measure was validated against objectively assessed sedentary time and also demonstrated reliability over repeated assessments. Previous studies have suggested questions that focus on TV viewing have demonstrated the strongest reliability and validity among non-occupational sedentary behaviour questions. It was not possible to adjust models for diet, as such data were not collected. However, recent data has suggested that snack food and TV viewing were independently associated with metabolic syndrome. The biomarkers were only collected once at baseline although variables such as CRP have demonstrated acceptable stability over long term follow-up. Our follow-up period was relatively short and reverse causation cannot be ruled out, although in sensitivity analyses the results remained unchanged after removal of deaths in the first year of follow up. Despite the relatively short follow-up it is likely that TV viewing reflects a habitual behaviour across the lifecourse.
may be associated with certain types of cancers,\textsuperscript{34} although we were unable to sub-type specific cancers owing to lack of events.

In conclusion, the association between TV viewing and mortality was partly mediated by inflammatory markers, although the relationship remains largely unexplained.
Conflict of interest

None of the authors have any competing interests to declare.

Funding sources

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Author contributions

Hamer had full access to the data, and takes responsibility for the integrity and accuracy of the results. Hamer drafted the paper, performed analyses and designed the study. Yates and Demakakos contributed to the concept and design of the study and critical revision of the manuscript.
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Table 1. Descriptive characteristics of the sample at baseline.

<table>
<thead>
<tr>
<th></th>
<th>&lt;2 hrs/d (n=877)</th>
<th>2 &lt; 4 hrs/d (n=2833)</th>
<th>4 &lt; 6 hrs/d (n=2279)</th>
<th>≥ 6 hrs/d (n=2462)</th>
<th>p-value&lt;sup&gt;c&lt;/sup&gt;</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (mean [SD] yrs)</td>
<td>63.3±10.1</td>
<td>64.1±9.9</td>
<td>65.6±9.8</td>
<td>65.4±9.9</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Men (%)</td>
<td>51.0</td>
<td>46.8</td>
<td>42.5</td>
<td>40.9</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Current smokers (%)</td>
<td>6.4</td>
<td>10.9</td>
<td>12.6</td>
<td>18.8</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Physically active&lt;sup&gt;a&lt;/sup&gt; (%)</td>
<td>44.0</td>
<td>37.2</td>
<td>27.6</td>
<td>22.1</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Regular alcohol intake&lt;sup&gt;b&lt;/sup&gt; (%)</td>
<td>70.0</td>
<td>68.4</td>
<td>58.4</td>
<td>54.2</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>At least one chronic illness (%)</td>
<td>45.6</td>
<td>47.8</td>
<td>55.7</td>
<td>59.5</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Depression (%)</td>
<td>12.0</td>
<td>9.7</td>
<td>12.6</td>
<td>18.5</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Disability (%)</td>
<td>18.1</td>
<td>18.4</td>
<td>24.8</td>
<td>33.7</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Obesity (%) (BMI ≥ 30 kg/m&lt;sup&gt;2&lt;/sup&gt;)</td>
<td>17.3</td>
<td>26.3</td>
<td>34.2</td>
<td>38.3</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>HDL-cholesterol (mmol/l)</td>
<td>1.66±0.45</td>
<td>1.58±0.42</td>
<td>1.56±0.41</td>
<td>1.52±0.40</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Triglycerides (mmol/l)</td>
<td>1.52±0.88</td>
<td>1.67±0.96</td>
<td>1.72±0.93</td>
<td>1.78±1.00</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>HbA1C (%)</td>
<td>5.76±0.42</td>
<td>5.81±0.62</td>
<td>5.89±0.68</td>
<td>5.95±0.77</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Log C-reactive protein (mg/l)</td>
<td>0.95±0.66</td>
<td>1.08±0.68</td>
<td>1.24±0.74</td>
<td>1.28±0.74</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Fibrinogen (g/l)</td>
<td>3.21±0.54</td>
<td>3.29±0.52</td>
<td>3.40±0.54</td>
<td>3.43±0.57</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Systolic Blood Pressure (mmHg)</td>
<td>129 ± 16</td>
<td>132 ± 17</td>
<td>133 ± 17</td>
<td>134 ± 18</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Prevalent CVD</td>
<td>22.2</td>
<td>23.9</td>
<td>27.7</td>
<td>31.1</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Prevalent cancer</td>
<td>7.4</td>
<td>6.8</td>
<td>7.9</td>
<td>7.1</td>
<td>0.40</td>
</tr>
</tbody>
</table>

Continuous variables are presented as mean ± SD and categorical variables as %.
<sup>a</sup> defined as vigorous intensity activity at least once per week.
<sup>b</sup> defined as alcohol intake at least once per week.
<sup>c</sup> p-values reflect differences across groups, examined using Chi-squared tests for categorical variables and ANOVA for continuous data.
Table 2. Cox proportional Hazards regression for association between TV viewing and mortality 
(n=8,451)

<table>
<thead>
<tr>
<th>TV exposure</th>
<th>Deaths/N</th>
<th>Model 1 HR (95% CI)</th>
<th>Model 2 HR (95% CI)</th>
<th>Model 3 HR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>&lt; 2hr/d</td>
<td>21/877</td>
<td>1.0 (Ref)</td>
<td>1.0</td>
<td>1.0</td>
</tr>
<tr>
<td>2&lt; 4 hr/d</td>
<td>104/2833</td>
<td>1.55 (0.97, 2.48)</td>
<td>1.63 (1.02, 2.61)</td>
<td>1.64 (0.96, 2.81)</td>
</tr>
<tr>
<td>4 &lt;6 hr/d</td>
<td>95/2279</td>
<td>1.62 (1.01, 2.59)</td>
<td>1.49 (0.92, 2.39)</td>
<td>1.66 (0.97, 2.84)</td>
</tr>
<tr>
<td>≥ 6 hr/d</td>
<td>150/2462</td>
<td>2.49 (1.58, 3.93)</td>
<td>1.98 (1.25, 3.15)</td>
<td>2.00 (1.18, 3.39)</td>
</tr>
<tr>
<td>Per SD increase$^a$</td>
<td>1.26 (1.15, 1.38)</td>
<td>1.17 (1.06, 1.28)</td>
<td>1.18 (1.05, 1.33)</td>
<td></td>
</tr>
</tbody>
</table>

Model 1; adjusted for age and sex.  
Model 2; adjusted for age, sex, physical activity, smoking, alcohol intake, depressive symptoms, long standing illness, disability (ADLs/IADLs)  
Model 3; adjusted for same covariates in model 2 after removal of 72 deaths in first year of follow up.  
$^a$Treated as a standardised score (1 SD=4.2 hrs/d TV viewing)
Table 3. Cox proportional Hazards regression for association between biological risk factors and mortality (n=5,033; 149 deaths).

<table>
<thead>
<tr>
<th>Risk factor</th>
<th>Model 1 HR per SD(^a) (95% CI)</th>
<th>Model 2 HR per SD (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>HDL-Cholesterol</td>
<td>0.75 (0.62, 0.91)</td>
<td>0.82 (0.66, 1.02)</td>
</tr>
<tr>
<td>HbA1C</td>
<td>1.14 (0.98, 1.32)</td>
<td>1.04 (0.89, 1.23)</td>
</tr>
<tr>
<td>Triglycerides</td>
<td>1.08 (0.91, 1.28)</td>
<td>1.03 (0.86, 1.25)</td>
</tr>
<tr>
<td>Log C-reactive protein</td>
<td>1.34 (1.22, 1.46)</td>
<td>1.13 (0.99, 1.29)</td>
</tr>
<tr>
<td>Fibrinogen</td>
<td>1.60 (1.39, 1.85)</td>
<td>1.40 (1.16, 1.68)</td>
</tr>
</tbody>
</table>

Model 1 adjusted for age, sex, BMI
Model 2 adjusted for age, sex, BMI, and mutually for other risk factors presented.
\(^a\)Standard deviation (SD) for each risk factor was: HDL-Cholesterol (0.42 mmol/l); HbA1C (0.69%); Triglycerides (0.96 mmol/l); log C-reactive protein (0.27 mg/l); Fibrinogen (0.55 g/l)
Table 4. Biological mediation analyses containing sub-sample from nurse visit (n=5,033)

<table>
<thead>
<tr>
<th></th>
<th>Basic model&lt;sup&gt;a&lt;/sup&gt;</th>
<th>Basic + inflammatory&lt;sup&gt;b&lt;/sup&gt;</th>
<th>Basic + metabolic&lt;sup&gt;c&lt;/sup&gt;</th>
<th>Final model&lt;sup&gt;d&lt;/sup&gt;</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>All cause mortality</strong>&lt;br&gt;(149 events)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>HR per SD unit&lt;sup&gt;e&lt;/sup&gt;</td>
<td>1.24 (1.07, 1.44)</td>
<td>1.20 (1.04, 1.39)</td>
<td>1.24 (1.07, 1.44)</td>
<td>1.21 (1.05, 1.41)</td>
</tr>
<tr>
<td>In HR</td>
<td>0.217</td>
<td>0.183</td>
<td>0.218</td>
<td>0.194</td>
</tr>
<tr>
<td><strong>CVD mortality</strong>&lt;br&gt;(29 events)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>HR per SD unit&lt;sup&gt;e&lt;/sup&gt;</td>
<td>1.50 (1.11, 2.03)</td>
<td>1.44 (1.06, 1.95)</td>
<td>1.51 (1.11, 2.05)</td>
<td>1.43 (1.06, 1.95)</td>
</tr>
<tr>
<td>In HR</td>
<td>0.404</td>
<td>0.361</td>
<td>0.412</td>
<td>0.361</td>
</tr>
</tbody>
</table>

<sup>a</sup> adjusted for age, sex, physical activity, smoking, alcohol intake, depressive symptoms, long standing illness, disability (ADLs/IADLs), body mass index, systolic blood pressure.

<sup>b</sup> Basic model (as above) with additional adjustment for log C-reactive protein and fibrinogen.

<sup>c</sup> Basic model with additional adjustment for HbA1C, HDL-C, triglycerides.

<sup>d</sup> Final model adjusted for age, sex, physical activity, smoking, alcohol intake, depressive symptoms, long standing illness, disability (ADLs/IADLs), body mass index, systolic blood pressure, log C-reactive protein, fibrinogen, HbA1C, HDL-C, triglycerides.

<sup>e</sup> Hazard ratio (HR) for TV viewing modelled continuously per SD increase (4.2 hrs/d)