Experimental study of workpiece-level variability in blind-via electroplating

This item was submitted to Loughborough University's Institutional Repository by the/an author.


Additional Information:

• This is an article from the journal, Proceedings of the IMechE, Part B: Journal of Engineering Manufacture [© IMechE]. It is also available at: http://journals.pepublishing.com/content/119784/?p=483cf2f36594441edbed851d8f3b10493&pi=0

Metadata Record: https://dspace.lboro.ac.uk/2134/4751

Version: Published

Publisher: Professional Engineering Publishing / © IMechE

Please cite the published version.
This item was submitted to Loughborough’s Institutional Repository (https://dspace.lboro.ac.uk/) by the author and is made available under the following Creative Commons Licence conditions.

For the full text of this licence, please go to: http://creativecommons.org/licenses/by-nc-nd/2.5/
Experimental study of workpiece-level variability in blind-via electroplating

G K K Poon1*, J S F Chan2 and D J Williams3
1Department of Manufacturing Engineering and Engineering Management, City University of Hong Kong, Hong Kong
2Engineering Mechanics, Materials and Design, School of Science and Technology, Open University of Hong Kong, Hong Kong
3Wolfson School of Mechanical and Manufacturing Engineering, Loughborough University, Loughborough, Leicestershire, UK

Abstract: The acid copper electroplating process for the manufacture of printed wire boards was studied by statistical techniques. The objectives of this study were to investigate the effects of process and product parameters on the workpiece-level uniformity during the acid copper plating of blind vias and to explore the minimization of the deposit thickness variation. The parameters studied were the concentrations of copper sulphate, sulphuric acid and additive, average current density (ACD), electrode separation (ES), the aspect ratio and the depth ratio of the via holes. Multifactor two-level factorial and the central composite rotatable five-level experiments were designed and conducted sequentially to generate statistical process models. Only the average current density and the electrode separation were found to be significant. A second-order model was then developed for the process in the proximity of the optimum region and verified experimentally to locate the optimum combinations of ACD and ES with respect to minimum thickness variability across the whole workpiece. Post-optimal analysis showed that the optimum solution was more sensitive to the electrode separation than the average current density.

Keywords: blind vias, acid copper electroplating, statistical experimentation, variability control

NOTATION

A  concentration of sulphuric acid (g/l)
B  concentration of copper sulphate (g/l)
C  average current density (mA/cm²)
D  electrode separation (cm)
E  interhole distance (cm)
F  corrected aspect ratio (mm)
G  depth ratio
L  thickness of a printed circuit board (mm)
R  radius of a via hole (mm)
SD  standard deviation of plating thickness along the via hole wall, as predicted by the statistical models (μm)
SD_min  minimum standard deviation of plating thickness along the via hole wall, as predicted by the statistical models (μm)

1 INTRODUCTION

The plated through-hole (PTH) and blind via (BV) are two widely used means of interconnection between the wiring layers of multilayer printed wiring boards (PWBs). In particular, the BV is regarded as the most efficient method of reducing the volume of space consumed by the layer-to-layer interconnects, and its importance is growing with the impending drive toward smaller geometry PWBs and integrated circuit package substrates [1]. The significance of uniform plating thickness of PTHs over the whole workpiece (PWB), discussed by Kessler and Alkire [2], applies equally to the electroplating of BVs. Reviews of previous work in the current distribution modelling of electrodeposition processes [3, 4] show that researchers have been concentrating on the feature-level plating uniformity of PTHs, i.e. the thickness distribution along the wall of a particular through-hole (see, for example, references [5] to [7]). Based on fundamental electrochemical principles, these models simulated the effect of various product and process parameters on the current...
distribution under well-defined assumptions and polarization characteristics. However, their proposed solution algorithms using numerical techniques, such as finite element and boundary element methods, are not readily extendable to the workpiece-level thickness distribution, i.e. the distribution of wall thickness among several through-holes over a substrate surface. Little modelling work has been done for the electroplating of BVs, at both the feature and the workpiece level.

On the other hand, the multiparameter nature of the electrodeposition process lends itself to modern methods of experimental design, but little has been published in the literature on this approach to solving problems. Empirical modelling of the through-hole plating process with statistically designed experiments was attempted by some authors, but only at the level of features [8–11], and no similar work was reported for the plating of BVs. Generally, these activities represent successful applications of simple designed experiments to obtain obvious directions of process improvement, usually in the form of a set of recommended factor settings. Empirical modelling and optimization of the process, however, can rarely be achieved and verified in a single set of experimental trials. Different experimental techniques are needed to achieve specific purposes during the step-by-step modelling process. This paper presents how planned matrix experiments, namely fractional factorials and central composite designs, can be conducted sequentially to model empirically the main effects as well as possible interactions of major process parameters on the average plating thickness and its variations in a BV electroplating process. This approach allows the optimal combination of these factors to be systematically obtained to minimize the workpiece-level plating thickness variability within the allowable process windows.

2 EXPERIMENTAL DETAILS

A computer numerically controlled (CNC) drilling machine was used to drill the BVs in the test boards using part programs (GM codes) developed with a text editor. Four layouts of the test boards are shown in Fig. 1 to illustrate the combinations of the various levels of interhole distance and the corrected aspect ratio. (The corrected aspect ratio, $L^2/R$, suggested in references [2] and [12] to [14] as a better indicator of the difficulty in achieving uniform plating of through-holes, is used here instead of the conventional definition of aspect ratio $L/R$. $L$ is the thickness of the printed circuit board (PCB) and $R$ is the radius of the via hole.) Depending on the dimension of the interhole distance (see Table 1), a maximum of 30 and a minimum of eight BVs were drilled in the test boards.

Before the copper plating process was performed, the non-conductive interior of the BVs was first metallized with a thin film of copper ($\approx 0.5 \mu m$) by the electroless copper process. Electroless solutions were prepared using analytical grade commercial reagents [15] and deionized water.

The plating bath employed was an acid copper sulphate solution, widely used in the PCB industry, containing copper(II) sulphate ($CuSO_4 \cdot 5H_2O$), sulphuric acid ($H_2SO_4$), chloride ions ($Cl^-$) and brightener (a proprietary organic additive [16]). Copper sulphate was the source of the metal (copper) deposit while the acid enhanced the solution conductivity. Sufficient chloride ions (60–80 ppm) and brightener (10 ppm) were included to give a bright and fine-grained plating. The anodes used were phosphorized copper plates. All the plating was performed in an air-agitated purpose-built testing cell which allowed precise control of the interelectrode separation and bath chemistry. The plating duration

---

![Fig. 1 Test boards of the BV experiments (not to scale)](image-url)
was fixed at 20 min for all the experimental runs, after which the deposit thickness along the hole walls was measured by an X-ray coating thickness gauge on sectioning.

The parameters studied in this work included the concentration of copper sulphate, sulphuric acid, average current density, electrode separation, corrected aspect ratio, interhole distance and the depth ratio of the via holes. The experimental strategies and designs are discussed in the next section.

3 MATRIX DESIGN AND RESULTS ANALYSIS

3.1 Experimental strategy

A sequential experimental strategy was adopted in this research. Two-level fractional factorial design (FFD) [17] was first used to screen out the significant parameters and to obtain first-order models of their effects on the deposit thickness variability. The model indicated the influence of these parameters, and their interactions (if any) on the workpiece-level variability within the process window, as well as the direction of optimizing the performance measure. A five-level central composite design [18] was then used to obtain a second-order model and to construct a response surface for the optimization of the plating process in order to minimize the thickness variability. Two-factor experiments were performed as complements to the factorial designs to verify the adequacy of the model built and to help the interpretation of parameters’ effects and the behaviour of the optimum solutions.

3.2 Factor screening

Table 1 summarizes the process and product parameters studied in the first stage of experimentation and their levels. The ranges of A, B, C and D cover the usual operating window while those of E and F represent the limits of the configuration of the purpose-built testing cell and the thickness of the test board.

A 32-run $2^{17}$ FFD [17] was used in the BV screening experiments. This design enabled the simultaneous study of the main effects of all the parameters and their possible interactions in a small number of runs. The resolution VI design was constructed using a $2^5$ full factorial as the first five columns and the defining relationship $I = ABCDF = ABDEG$ for the last two columns. Such high resolution ensured that each main effect was aliased only with high-order interactions, e.g. three- or four-factor interactions, which could be assumed to be negligible. The mean and standard deviation (SD) of the plating thickness readings taken for each board were used as the responses. All the hole walls were measured but the largest and smallest thickness values were excluded from the calculations to reduce the bias caused by extreme data. The parameter settings and the responses of the screening experiments, which were conducted in a randomized order, are summarized as Table 2. All the effect calculations and analysis were carried out using JMP Release 3.2 while the half-normal probability and response surface plots were

<table>
<thead>
<tr>
<th>Factors</th>
<th>Levels</th>
</tr>
</thead>
<tbody>
<tr>
<td>A</td>
<td></td>
</tr>
<tr>
<td>B</td>
<td></td>
</tr>
<tr>
<td>C</td>
<td></td>
</tr>
<tr>
<td>D</td>
<td></td>
</tr>
<tr>
<td>E</td>
<td></td>
</tr>
<tr>
<td>F</td>
<td></td>
</tr>
<tr>
<td>G</td>
<td></td>
</tr>
</tbody>
</table>

Table 1 Factor settings of the BV screening experiments

<table>
<thead>
<tr>
<th>Factors</th>
<th>Levels</th>
</tr>
</thead>
<tbody>
<tr>
<td>A</td>
<td></td>
</tr>
<tr>
<td>B</td>
<td></td>
</tr>
<tr>
<td>C</td>
<td></td>
</tr>
<tr>
<td>D</td>
<td></td>
</tr>
<tr>
<td>E</td>
<td></td>
</tr>
<tr>
<td>F</td>
<td></td>
</tr>
<tr>
<td>G</td>
<td></td>
</tr>
</tbody>
</table>

Table 2 Parameter settings and the responses of the BV screening experiment

EXPERIMENTAL STUDY OF WORKPIECE-LEVEL VariABILITY 523

constructed by STATISTICA Release 5.0 and MATLAB Release 4.2.

3.3 Analysis of the mean response

The half-normal probability plot [19] of the mean responses, Fig. 2a, shows that the significant parameters are the average current density (C), electrode separation (D) and the two two-factor interactions: sulphuric acid × copper sulphate (AB) and average current density × electrode separation (CD). An ANOVA test confirmed also their significance ($\alpha < 0.005$). Figure 2b illustrates the strong positive effects of C, D, AB and CD on the mean plating thickness of the BVs and the corresponding interaction plots are given in Fig. 3.

The mean response reflects the average plating rate under the particular plating condition. Figure 3a shows that maximum average plating rate can be achieved when the concentrations of both sulphuric acid (A) and copper sulphate (B) are at their high (+) level, i.e. 200 and 90 g/l respectively. However, this maximum plating rate is found to be not significantly different from that obtained when both A and B are at their low (−) level, i.e. 165 and 65 g/l respectively. A first-order model was built for the BV process using the significant parameters identified above, with the ANOVA tests performed by pooling the residual sum of squares to give an estimate of the random errors for the experiments. The three-term BV model ($\alpha < 0.0001$) is found to be highly significant and given by the following equation in coded form:

$$SD = 1.7830 + 0.3920C + 0.2165D$$

(1)

where SD is the predicted standard deviation and C and D take the values −1 or +1 for the low and high levels of the corresponding parameters. Analysis of the residuals

without a proportional change of the other one will result in a lower average plating rate. On the other hand, Fig. 3b shows that maximum average plating rate results when high average current density (43.04 mA/cm$^2$) and large electrode separation (17.78 cm) are employed. While the first condition is common to all Faradaic reactions, the interaction plot shows that this Faradaic effect is attenuated by the magnitude of the electrode separation, which manifests itself as the strong interaction between the two parameters.

3.4 First-order process model of SD responses

The half-normal probability plot of Fig. 4 shows that the average current density (C) and the electrode separation (D) are the only two active parameters affecting the workpiece-level uniformity of the BVs. Both of the two parameters have strong positive effect on the thickness variation, as illustrated in the effect plots of Fig. 5, which means that minimum variability results when both of them are at their low level, i.e. 32.28 mA/cm$^2$ and 12.7 cm respectively. A first-order model was built for the BV process using the significant parameters identified above, with the ANOVA tests performed by pooling the residual sum of squares to give an estimate of the random errors for the experiments. The three-term BV model ($\alpha < 0.0001$) is found to be highly significant and given by the following equation in coded form:

$$SD = 1.7830 + 0.3920C + 0.2165D$$
Fig. 3  Interaction plots for the parameters A, B, C and D for the mean responses

Fig. 4  Half-normal probability plot of the BV screening experiment (SD responses)

Fig. 5  Analysis of mean plot for the SD responses
of the first-order model revealed no abnormality of concern. The lack of fit, which is found to be insignificant ($\hat{c}_1 = 0.469$), shows that the first-order models accurately describe the response surface and can be used to search for levels of the parameters that would produce more optimal values of the response, i.e. smaller thickness variation (SD). The response surface of the first-order model is shown in Fig. 6, which indicates clearly the location of minimum variability, $[C(-), D(-)]$, within the experimental range. This corner represents also the approximate location of the centre point of the CCD design for the second-order model building.

### 3.5 Second-order model building

Four more experiments were performed in the proximity of the minimum variability, $[C(-), D(-)]$, identified in the first-order model above. The factor settings and responses are summarized in Table 3. All the other insignificant parameters were held constant at their low (−) levels. The same measurement and response scheme was adopted in this round of experiments and the average plating efficiency was found to be 87.3 per cent.

The results show that further reduction of thickness variation can be achieved within the region and the combination of $C = 35.0 \text{mA/cm}^2$ and $D = 14.0 \text{cm}$ was selected as the centre point of the central composite design for second-order model building.

Using the centre point identified above, a 10-run rotatable central composite design [17] involving the parameters $C$ and $D$ was designed and conducted to build a second-order model for the workpiece-level variation. Table 4 summarizes the factor levels and setting of the central composite design employed.

The standard deviations of the hole–wall plating thickness were taken as the responses and are

---

**Table 3** Design points and responses in the proximity of $[C(-), D(-)]$

<table>
<thead>
<tr>
<th>Board</th>
<th>Average current density (mA/cm$^2$)</th>
<th>Electrode separation (cm)</th>
<th>Standard deviation (µm)</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>35.5</td>
<td>14.0</td>
<td>1.103</td>
</tr>
<tr>
<td>2</td>
<td>32.3</td>
<td>14.0</td>
<td>0.894</td>
</tr>
<tr>
<td>3</td>
<td>34.4</td>
<td>13.3</td>
<td>0.820</td>
</tr>
<tr>
<td>4</td>
<td>32.3</td>
<td>12.7</td>
<td>0.953</td>
</tr>
</tbody>
</table>

---

**Table 4** Factor levels and settings of the BV CCD experiment

<table>
<thead>
<tr>
<th>Factors</th>
<th>Fixed level</th>
<th>Factor levels</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>$-1.414$</td>
</tr>
<tr>
<td>A Sulphuric acid (g/l)</td>
<td>180</td>
<td>—</td>
</tr>
<tr>
<td>B Copper sulphate (g/l)</td>
<td>75</td>
<td>—</td>
</tr>
<tr>
<td>C Average current density (mA/cm$^2$)</td>
<td>—</td>
<td>31.2</td>
</tr>
<tr>
<td>D Electrode separation (cm)</td>
<td>—</td>
<td>12.2</td>
</tr>
<tr>
<td>E Interhole distance (cm)</td>
<td>0.45</td>
<td>—</td>
</tr>
<tr>
<td>F Corrected aspect ratio (mm)</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>G Depth ratio</td>
<td>0.875</td>
<td>—</td>
</tr>
</tbody>
</table>
summarized in Table 5. Twenty thickness readings were taken from the 10 vias of each test board and the largest and smallest readings were discarded in the calculation. An average plating efficiency of 84.5 per cent was recorded in this round.

The six-term full second order model in coded form is

\[
SD = 1.6573 - 0.0982C + 0.3667D - 0.1887C^2 \\
+ 0.7386D^2 + 0.5537CD
\]

(2)

The ANOVA test confirmed the significance of the model \( (\alpha = 0.0045) \); its \( R^2_A \) (0.9266) performance is better than that of the first-order model derived above although its \( R^2_{\text{prediction}} \) (0.7781) is slightly lower. The full model is also the best-fitted model under the backward elimination process \([20]\) and Mallow’s criteria \([21]\). Residual analysis of the model revealed no abnormality of concern. Figure 7 shows that the response surface of the second-order mode is a saddle system \([18]\), in which the stationary point is neither a maximum nor a minimum point.

The location of the minimum variation within the experimental space can be found on the plane of \( C = 1.414 \), on which the second-order model becomes

\[
SD = 0.7386D^2 + 1.1496D + 1.1410
\]

(3)

Differentiating equation (3) with respect to \( D \) yields the stationary point of \( C = 1.414 \) (36 mA/cm²) and \( D = -0.7782 \) (12.98 cm), which gives a minimum standard deviation, \( SD_{\text{min}} \), of 0.6934 \( \mu \)m. The locus of the minimum standard deviation is given in Fig. 8 while Figs 9a and b show respectively the quadratic relationship between \( SD_{\text{min}} \) and \( C \) and the linear relationship between the optimum values of \( C \) and \( D \). Figures 8 and 9a show that \( SD_{\text{min}} \) is smaller at the two extreme values of \( C \) and \( D \). Figure 9b shows that the sensitivity of the optimum solution to the electrode separation is 2.7 times higher than to the average current density. Physically, this means that, as long as the process is operating along the path shown in Fig. 9b, the workpiece-level variability of the BVs will be minimized with respect to that particular combination of average current density and electrode separation, and in all cases is limited to the range 0.7–1.7 \( \mu \)m. These correspond to a 29–71 per cent reduction of variability when compared with the worst scenario (2.4 \( \mu \)m) of the first-order response surface (Fig. 6).

### 3.6 Model verification

Twenty-four more experiments were performed to verify the adequacy of the second-order model in predicting the workpiece-level plating uniformity within the experimental range. The standard deviation of 20 hole–wall plating thickness readings was determined from each test board.
Two replicates were run for each of the twelve combinations of C and D, while the mean SDs were taken as the responses. The factor settings and responses are summarized in Table 6. Predicted responses of the second-order model [equation (2)] are plotted against the actual standard deviations obtained from the model building (Table 5) and verification runs (Table 6) in Fig. 10. It shows that the full second-order model is able to predict 92.6 per cent of the variations, which is much higher than the expected performance of 77.81 per cent suggested by $R^2_{\text{prediction}}$ of the model determined in Section 3.5.

**Table 6** Factor settings and responses of the BV verification experiment

<table>
<thead>
<tr>
<th>Electrode separation (cm)</th>
<th>Average current density (mA/cm²)</th>
<th>SD (μm)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Actual</td>
<td>Predicted</td>
</tr>
<tr>
<td>12.7</td>
<td>31.0</td>
<td>2.369</td>
</tr>
<tr>
<td></td>
<td>32.0</td>
<td>2.0194</td>
</tr>
<tr>
<td></td>
<td>35.5</td>
<td>1.2117</td>
</tr>
<tr>
<td>14.0</td>
<td>30.0</td>
<td>1.2637</td>
</tr>
<tr>
<td></td>
<td>33.0</td>
<td>1.7062</td>
</tr>
<tr>
<td></td>
<td>34.0</td>
<td>1.4435</td>
</tr>
<tr>
<td></td>
<td>35.0</td>
<td>1.1713</td>
</tr>
<tr>
<td>15.2</td>
<td>30.5</td>
<td>2.3076</td>
</tr>
<tr>
<td></td>
<td>31.5</td>
<td>2.5554</td>
</tr>
<tr>
<td></td>
<td>34.0</td>
<td>2.8616</td>
</tr>
<tr>
<td></td>
<td>37.0</td>
<td>3.1164</td>
</tr>
<tr>
<td></td>
<td>39.0</td>
<td>2.8862</td>
</tr>
</tbody>
</table>

Two replicates were run for each of the twelve combinations of C and D, while the mean SDs were taken as the responses. The factor settings and responses are summarized in Table 6. Predicted responses of the second-order model [equation (2)] are plotted against the actual standard deviations obtained from the model building (Table 5) and verification runs (Table 6) in Fig. 10. It shows that the full second-order model is able to predict 92.6 per cent of the variations, which is much higher than the expected performance of 77.81 per cent suggested by $R^2_{\text{prediction}}$ of the model determined in Section 3.5.

**4 DISCUSSION**

Previous work on the modelling of PTH plating shows that the feature scale uniformity is limited by the ohmic resistance (secondary current density effect) rather than the mass transfer resistance (cupric ion concentration, agitation effect) of the system [6, 14, 22]. Similar results are not found in the literature for BV plating. The statistical models developed here show that the ohmic limitation applies also to the workpiece-level thickness variation for BV plating, although the mass transfer limitation has more influence on optimum BV plating than the PTH plating, as reflected in the mean response analysis of Section 3.3.

The electrode separation was found to be significant in affecting the workpiece-level variation of the BVs, and the uniformity deteriorated with the electrode separation. This is evident from both the first- and the second-order models, equations (1) and (2) respectively, of the process. It is believed that, under the same agitation scheme, the convective flow of electrolyte into the BVs is enhanced when the electrode separation is reduced and hence the deposit uniformity is improved. The first-order model shows also that the workpiece-level variability is minimized when both the average current density and the electrode separation are at their low levels. Such a combination, however, will result in a lower average plating rate (Fig. 3b). Analysis of the mean thickness responses of the screening experiments shows that the interaction effect between the sulphuric acid and copper sulphate (AB) was significant in affecting the average plating rate. The plating rate was maximized when a high acid concentration was coupled with a high sulphate concentration, or its reverse, as shown in Fig. 3a. This can be explained by the fact that addition of copper sulphate has a negative effect on the acidity and hence the conductivity of the acid–sulphate system [23], while conductivity has a negative effect on the reaction rate of plating without convection [24], with which the BV plating process is comparable. Therefore, the negative effect of the acid...
on the plating rate was counterbalanced by the copper sulphate when these two components were in appropriate proportion. Such an interaction effect, which does not affect the thickness variation, could be used to obtain a compromise between the two conflicting requirements of minimum workpiece-level variability and high average plating rate.

5 CONCLUSION

This paper summarizes the sequential experimental studies of the BV copper plating process. Significant process parameters affecting the workpiece-level uniformity were identified as the average current density and the electrode separation. Statistically proven and electrochemically informed models were developed for the copper plating of BVs to predict and minimize, within the operability range and experimental accuracy, the workpiece-level thickness variability. As these models are defined by parameters independent of the physicochemical and polarization properties of the electrolytic system, they can be used for the variability control and optimization of industrial BV plating processes.

REFERENCES

15 Alfachimici. Technical Bulletins 8506 (Cuprolite X-84), 8510 (Microincide 1207), 8000 (Uniphase MLX A), 8501 (Uniphase MLX A+B), 8508 (Drag-Stop 79), 1985.


